

Isaac's probability function, v.6

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Sampling a probability-density

Given a prob-density $D()$, how do we produce a D -random number?

The corresponding CDF (Cumulative Distribution Fnc) is

$$F(x) := \int_{-\infty}^x D \stackrel{\text{specifically}}{\int_{(-\infty, x]}} D .$$

So for each \mathbf{x}_0 real:

*: A D -random point $q \in \mathbb{R}$ has the property that $P_D(q \leq \mathbf{x}_0)$ equals $F(\mathbf{x}_0)$.

Suppose $D > 0$ on \mathbb{R} . Then F is continuous and strictly-increasing, mapping $\mathbb{R} \leftrightarrow (0, 1)$. Function-inverse F^{-1} maps $(0, 1)$ onto \mathbb{R} .

Fix $\mathbf{y}_0 \in (0, 1)$. Then \mathbf{y}_0 is the probability –using the uniform-distribution on $(0, 1)$ – that a chosen point y in $(0, 1)$ satisfies $y \leq \mathbf{y}_0$.

With $q := F^{-1}(y)$ and $\mathbf{x}_0 := F^{-1}(\mathbf{y}_0)$, note

$$y \leq \mathbf{y}_0 \iff q \leq \mathbf{x}_0 ,$$

since F^{-1} is strictly increasing. Hence the probability that $q \leq \mathbf{x}_0$ equals $\mathbf{y}_0 \stackrel{\text{note}}{=} F(\mathbf{x}_0)$, as desired in (*).

1a: ALGORITHM A: Use the uniform-distribution to pick $y \in (0, 1)$, then **Return $F^{-1}(y)$** . \square

Possible alternative. Suppose density $D()$ is zero outside of some interval $[L, H]$.

If F^{-1} is easy to compute, then use (1a). However, if F^{-1} is difficult [or if the density is zero on intervals inside $[L, H]$, so there is no global F^{-1}], one can generate a random number directly from D , with the following *not-terribly-efficient* algorithm.

Let **Max** here mean the maximum [technically, *supremum*] of $D(x)$ over all $x \in [L, H]$.

Fix any number $M \geq \text{Max}$.

1b: ALGORITHM B: Pick x uniformly in $[L, H]$ and y uniformly in $[0, M]$. If $y \leq D(x)$, then **Return x** ; else go back and pick new x, y . \square

For this algorithm, our D need not be normalized. E.g, replacing $D()$ by $\frac{1}{3}D()$ would not change the induced prob-measure. [But it would increase the expected number of iterations before an x is returned.]

The expected # of iterations can be reduced, by lowering M to equal **Max**.

Technical. A collection \mathfrak{C} of \mathbb{R} -subsets is a σ -algebra if \mathfrak{C} is sealed under complement and countable union. I.e.,

$$\begin{aligned} B \in \mathfrak{C} &\implies [\mathbb{R} \setminus B] \in \mathfrak{C} \quad \text{and} \\ B_j \in \mathfrak{C} &\implies \left[\bigcup_{j=1}^{\infty} B_j \right] \in \mathfrak{C} . \end{aligned}$$

The **Borel σ -algebra** is the smallest σ -algebra owning all the open intervals.

A **probability measure** on \mathbb{R} assigns a probability $P(B)$ to each Borel set B . Its CDF (Cumulative Distribution Fnc) is $F(x) := P((-\infty, x])$.

A probability measure is **atomic** if $\exists x \in \mathbb{R}$ with $P(\{x\})$ positive. A probability has a **density** (PDF) IFF it is non-atomic. Moreover, its density is the derivative of its CDF. [THEOREM: A CDF is differentiable “almost everywhere”, that is, off of a probability-zero set.] \square

THE PROBLEM. Fix positive reals $L < H$ (Low < High) and biases α, β . We seek a single-humped probability measure on \mathbb{R} which we condition on interval $[L, H]$, then emit a random $\#$. There is a *seller* who will (grumpily) accept as little as L for his MacGuffin, and a *buyer* who will pay as much as H ; both values secret.

The biases α, β are the fair-market-opinions of the *seller* and *buyer*, respectively, of MacGuffin's value.

[Originally, I misunderstood the biases, thinking that Isaac wanted **affine**-invariance w.r.t "The Data", L, H, α, β . However, fair-market values relate to \$0, so only scale-invariance is appropriate. We *could* force scale-invariance by dividing all the given data by, say, H , and then multiply the resulting random# by H , but we'll do it directly.]

Interval conditioning. We'll use "The Data" to produce a CDF F and density D , then condition them on interval $[L, H]$, producing \tilde{F} and \tilde{D} . More precisely, our \tilde{F} will stretch vertically outside of $[0, 1]$, but if we restrict our attention to $[L, H]$, then \tilde{F} is a CDF.

This *almost*-CDF corresponding to F is

$$\begin{aligned} \tilde{F} &:= \mathcal{A} \circ F, \text{ using affine } \mathcal{A}(z) := \frac{z - \ell}{m} \\ \text{2a: with } \ell &:= F(L) \text{ and } m := F(H) - F(L). \text{ So} \\ \tilde{F}^{-1} &:= F^{-1} \circ \mathcal{A}^{-1}, \text{ where } \mathcal{A}^{-1}(y) := \ell + my. \end{aligned}$$

As \tilde{F} is a vertically stretched CDF, it can take on values >1 or <0 . Nonetheless, ALGORITHM A applies to \tilde{F}^{-1} , since the input is a point in $(0, 1)$. The actual CDF is $C_{0,1} \circ \tilde{F}$, using a **cutoff function**

$$\text{2b: } C_{\text{Low,High}}(x) := \text{Min}(\text{Max}(x, \text{Low}), \text{High}),$$

for real numbers $\text{Low} \leq \text{High}$.

When F is differentiable, so is \tilde{F} .

$$\text{2c: Density } \tilde{D} \text{ is derivative } \tilde{F}', \text{ restricted to interval } [L, H].$$

The only part of \tilde{D} we would wish to graph is on $[L, H]$. Thus we *never need* to work with $C_{0,1} \circ \tilde{F}$ [the actual CDF] but only with \tilde{F}^{-1} and \tilde{F}' .

Tools. The classic *sigmoid* func is

$$\begin{aligned} \text{3a: } S(x) &:= \frac{e^x}{e^x + 1} \stackrel{\text{note}}{=} \frac{1}{1 + e^{-x}}, \text{ with} \\ S^{-1}(y) &= \log\left(\frac{y}{1-y}\right) \stackrel{\text{note}}{=} \log\left(\frac{1}{1-y} - 1\right), \end{aligned}$$

where \log is natural logarithm, \log_e . Our S is a strictly-incr map of \mathbb{R} onto $(0, 1)$. Under rotation by 180° , its graph is symmetric about the point $(0, \frac{1}{2})$.

For future reference, the sigmoid's derivative is

$$\text{3b: } S'(x) = \frac{e^x}{[e^x + 1]^2}.$$

The rotational symmetry of $S()$ follows from showing that $S'()$ is even. Indeed,

$$S'(-x) = \frac{e^{-x}}{[e^{-x} + 1]^2} \cdot \left[\frac{e^x}{e^x}\right]^2 = \frac{e^x}{[1 + e^x]^2} \stackrel{\text{note}}{=} S'(x).$$

Construction. Our CDF will have form $F(x) := S(\text{Multiplier} \cdot [x - \tau])$, where **translation** τ is the center of density D . [Center τ might be outside of $[L, H]$, if the fair-market values mostly agree, and are well outside $[L, H]$.] To guarantee that the CDF is scale-invariant, the specific form will be

4a:
$$F(x) := S\left(c \cdot \frac{x - \tau}{L + H}\right),$$

where **concentration** c is multiplicatively-invariant; multiplying L, H, α, β by a constant does not change c . Increasing the positive number c will concentrate the probability around center τ .

A mult-invariant number is **reliability**,

4b:
$$\rho := 1 - \frac{|\beta - \alpha|}{\beta + \alpha},$$

which is close to 1 when the buyer's and seller's fair-market opinions are relatively close.

Details. Varying with scale are averages

4c:
$$V_{\text{Fair}} := \frac{\alpha + \beta}{2} \quad \text{and} \quad V_{\text{Range}} := \frac{L + H}{2},$$

and weighted-average

4d:
$$\tau := V_{\text{Fair}} \cdot \hat{\rho} + V_{\text{Range}} \cdot [1 - \hat{\rho}], \quad \text{where } \hat{\rho}$$

is some positive, non-decreasing function of ρ .

E.g, $\hat{\rho}$ could be ρ^3 .

Concentration c should also be a *positive, non-decreasing* fnc of ρ . [Likely you want $c \geq 1$ always. Isaac, we experimented with something like $c := [0.2 + \rho] \cdot 20$.]

Coding. Reals $M \neq 0$ and T determine an affine map

5:
$$\mathcal{A}_{T,M}(z) := \frac{z - T}{M}, \quad \text{with } \mathcal{A}_{T,M}^{-1}(y) := T + My.$$

 and (constant) derivative $\mathcal{A}'_{T,M} = \frac{1}{M}$.

Decide your formulas for “dials” $\hat{\rho}$ and c . Then:

Program. Define τ by (4b,4c,4d), then let

6.1:
$$F := S \circ \mathcal{A}_{\tau,M} \quad \text{where} \quad M := \frac{L + H}{c}.$$

Our *almost*-CDF is now

6.2:
$$\tilde{F} := \mathcal{A}_{\ell,m} \circ F, \quad \text{where}$$

$$\ell := F(L) \quad \text{and} \quad m := F(H) - F(L).$$

ALGORITHM A: Pick y uniformly in $(0, 1)$. Using S^{-1} from (3a), compute $q := \mathcal{A}_{\tau,M}^{-1}(S^{-1}(\mathcal{A}_{\ell,m}^{-1}(y)))$. I.e,

6.3:
$$q := \tau + M \cdot S^{-1}(\ell + my)$$

is your random number. □

Graphing. Density \tilde{D} is the derivative of \tilde{F} . I.e,

6.4:
$$\tilde{D}(x) = \frac{1}{m} \cdot S'\left(\frac{x - \tau}{M}\right) \cdot \frac{1}{M}$$

on $[L, H]$, using S' from (3b). □

NB. You *might* be able to avoid overflow^{♥1} by using formulae $S(x) = \frac{1}{[1 + e^{-x}]}$ and $S'(x) = \frac{e^{-x}}{[e^{-x} + 1]^2}$.

Alternatively, you can replace S by a sigmoid-like fnc, e.g

$$\mathfrak{S}(x) := \frac{1}{2} + \frac{\arctan(x)}{\pi}. \quad \text{Thus}$$

 7:
$$\mathfrak{S}^{-1}(y) = \tan\left(\left[y - \frac{1}{2}\right] \cdot \pi\right) \quad \text{and}$$

$$\mathfrak{S}'(x) = \frac{1}{\pi \cdot [1 + x^2]}.$$

^{♥1}But you might have underflow. You can *guarantee* no overflow, by using that S' is even, hence $S'(x) = \frac{e^{-|x|}}{[e^{-|x|} + 1]^2}$.